# On a 2-class Polling Model with Class-dependent Reneging, Switchover Times, and Phase-type Service 

Kevin Granville \& Steve Drekic<br>Department of Statistics \& Actuarial Science<br>University of Waterloo<br>MAM $9 \quad$ June 28-30, 2016

On a 2-class

1 Introduction and Preliminaries

2 Determination of the Steady-state Probabilities

3 Determination of the Waiting Time Distribution

4 Numerical Analysis

5 Concluding Remarks

## Polling Models

On a 2-class Polling Model...

## Kevin

 Granville \& Steve Drekic
## Introduction

 and PreliminariesDetermination of the Waiting Time Distribution

- A typical polling model consists of multiple queues attended by a single server in cyclic order.
- Due to its wide use in the areas of public health systems, transportation, and communication and computer networks, polling models have drawn considerable attention over the past fifty years.


## Notable References

- M.A.A. Boon, "Polling Models: From Theory to Traffic Intersections". PhD Thesis, Eindhoven: Technische Universiteit Eindhoven, 190 pages, 2011.
■ H. Levy \& M. Sidi, "Polling systems: applications, modeling and optimization". IEEE Transactions on Communications, Vol. COM-38, No. 10, pp. 1750-1760, 1990.
- H. Takagi, "Queueing analysis of polling models". ACM Computing Surveys, Vol. 20, No. 1, pp. 5-28, 1988.
- V.M. Vishnevskii \& O.V. Semenova, "Mathematical methods to study the polling systems". Automation and Remote Control, Vol. 67, No. 2, pp. 173-220, 2006.
- S. Borst, O. Boxma \& H. Levy, "The use of service limits for efficient operation of multistation single-medium communication systems", IEEE/ACM Transactions on Networking, Vol. 3, No. 5, pp. 602-612, 1995.


## Proposed Queueing Model

| On a 2-class |
| :--- |
| Polling |
| Model... |

$\quad$ Kevin
Granville \&
Steve Drekic


■ Poisson arrivals with rates $\lambda_{i}$

- Continuous phase-type service times
- FCFS, $k_{i}$-limited service discipline

■ Exponential switchover times with rates $v_{i}$
■ Exponential reneging times with rates $\alpha_{i}$

- All distributions are independent
- Finite buffer sizes $b_{i}<\infty$


## Proposed Queueing Model

On a 2-class
Polling
Model...

## Kevin

 Granville \& Steve Drekic
## Introduction

 and Preliminaries■ Service times for class- $i$ customers, $i=1,2$, are assumed to follow a continuous phase-type distribution (of dimension $\mathrm{s}_{i}$ ), with probability density function of the form

$$
f_{i}(\omega)=\underline{\beta}_{i} \exp \left\{S_{i} \omega\right\} \underline{S}_{0, i}^{\prime}, \omega>0
$$

■ Initial probability row vector is $\underline{\beta}_{i}=\left(\beta_{i, 1}, \beta_{i, 2}, \ldots, \beta_{i, s_{i}}\right), \sum_{j=1}^{s_{i}} \beta_{i, j}=1$.

- $S_{i}$ is an $\mathrm{s}_{i} \times \mathrm{s}_{i}$ rate matrix and $\underline{S}_{0, i}^{\prime}=-S_{i} \underline{e}_{\mathrm{s}_{i}}^{\prime}$, where $\underline{e}_{\mathrm{s}_{i}}^{\prime}$ is a column vector of $s_{i}$ ones.


## Steady-state Probabilities

## On a 2-class

Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and Preliminaries

- For $i=1,2$, let $X_{i}$ represent the number of class- $i$ customers present in the system, so that $0 \leq X_{i} \leq b_{i}$.
- Our first objective is to determine

$$
\left\{P_{m, n} ; m=0,1, \ldots, b_{1}, n=0,1, \ldots, b_{2}\right\}
$$

where $P_{m, n}$ denotes the steady-state joint probability that $X_{1}=m$ and $X_{2}=n$.

- Define an associated quantity $\pi_{m, n, l, y}$ representing the steady-state joint probability that $X_{1}=m, X_{2}=n$, the server being in position $I$, and the current phase of service being $y$ (with $y=0$ indicating that the system is in switchover mode).


## Steady-state Probabilities

- Dependent on $m$ and $n$, component $/$ takes on the following values:

$$
\begin{array}{rl}
m=n=0 \Longrightarrow & I= \\
k_{1}+k_{2}+1, k_{1}+k_{2}+2, \\
m \neq 0 \text { and } n=0 \Longrightarrow &  \tag{1}\\
m=0 \text { and } n \neq 0 \Longrightarrow & 1,2, \ldots, k_{1}, k_{1}+k_{2}+1, k_{1}+k_{2}+2, \\
m \neq 0 \text { and } n \neq 0 \Longrightarrow & k_{1}+1, k_{1}+2, \ldots, k_{1}+k_{2}, k_{1}+k_{2}+1, \\
& k_{1}+k_{2}+2, \\
m & I=\begin{array}{l}
1,2, \ldots, k_{1}, k_{1}+1, k_{1}+2, \ldots, k_{1}+k_{2}, \\
\\
\\
k_{1}+k_{2}+1, k_{1}+k_{2}+2 .
\end{array}
\end{array}
$$

- When $I=1,2, \ldots, k_{1}$, the server is serving its $I^{t h}$ customer from the class-1 queue.
- When $I=k_{1}+1, k_{1}+2, \ldots, k_{1}+k_{2}$, the server is serving its $\left(I-k_{1}\right)^{t h}$ customer from the class-2 queue.
- When $I=k_{1}+k_{2}+i$, the server is conducting a switchover out of the class- $i$ queue, $i=1,2$.


## Steady-state Probabilities

On a 2-class
Polling
Model...
Kevin
Granville \&
Steve Drekic

Introduction
and
Preliminaries
Determination
of the
Steady-state
Probabilities
Determination
of the
Waiting Time
Distribution
Numerical
Analysis
Concluding
Remarks

- Similarly, component $y$ depends on I in the following way:

$$
\begin{align*}
I=1,2, \ldots, k_{1} & \Longrightarrow y=1,2, \ldots, s_{1}, \\
I=k_{1}+1, k_{1}+2, \ldots, k_{1}+k_{2} & \Longrightarrow y=1,2, \ldots, s_{2},  \tag{2}\\
I=k_{1}+k_{2}+1, k_{1}+k_{2}+2 & \Longrightarrow y=0 .
\end{align*}
$$

## Steady-state Probabilities

On a 2-class Polling Model...

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis
Concluding Remarks

- When $m=n=0$ (i.e., the queue is empty), the system can only be in one of two kinds of switchover modes (as there are no customers to serve in either queue) and so $P_{0,0}=\pi_{0,0, k_{1}+k_{2}+1,0}+\pi_{0,0, k_{1}+k_{2}+2,0}$.
- It follows that

$$
\begin{aligned}
P_{0, n} & =\sum_{l=k_{1}+1}^{k_{1}+k_{2}} \sum_{y=1}^{s_{2}} \pi_{0, n, l, y}+\sum_{l=k_{1}+k_{2}+1}^{k_{1}+k_{2}+2} \pi_{0, n, l, 0}, n \geq 1, \\
P_{m, 0} & =\sum_{l=1}^{k_{1}} \sum_{y=1}^{s_{1}} \pi_{m, 0, l, y}+\sum_{l=k_{1}+k_{2}+1}^{k_{1}+k_{2}+2} \pi_{m, 0, l, 0}, m \geq 1
\end{aligned}
$$

and

$$
P_{m, n}=\sum_{l=1}^{k_{1}} \sum_{y=1}^{s_{1}} \pi_{m, n, l, y}+\sum_{l=k_{1}+1}^{k_{1}+k_{2}} \sum_{y=1}^{s_{2}} \pi_{m, n, l, y}+\sum_{l=k_{1}+k_{2}+1}^{k_{1}+k_{2}+2} \pi_{m, n, l, 0}, m, n \geq 1
$$

## Steady-state Probabilities

## On a 2-class

Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis

- Define the $0^{t h}$ steady-state probability row vector to be $\underline{\pi}_{0}=\left(\underline{\pi}_{0,0}, \underline{\pi}_{0,1}, \ldots, \underline{\pi}_{0, b_{2}}\right)$, where $\underline{\pi}_{0,0}=\left(\pi_{0,0, k_{1}+k_{2}+1,0}, \pi_{0,0, k_{2}+k_{2}+2,0}\right)$, and $\underline{\pi}_{0, n}, n=1,2, \ldots, b_{2}$, is a row vector of size $z_{1}=k_{2} s_{2}+2$.
- For $m \geq 1$, the $m^{t h}$ steady-state probability row vector is defined as $\underline{\pi}_{m}=\left(\underline{\pi}_{m, 0}, \underline{\pi}_{m, 1}, \ldots, \underline{\pi}_{m, b_{2}}\right)$, where $\underline{\pi}_{m, 0}$ is a row vector of size $k_{1} s_{1}+2$ and $\underline{\pi}_{m, n}, n=1,2, \ldots, b_{2}$, is a row vector of size $z_{2}=k_{1} s_{1}+z_{1}$.
- Referring to $X_{1}$ as the level of the process, we remark that level 0 is comprised of $n_{1}=b_{2} z_{1}+2$ sub-levels, whereas each non-zero level consists of a total of $n_{2}=b_{2} z_{2}+k_{1} s_{1}+2$ sub-levels.


## Steady-state Probabilities

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical
Analysis
Concluding Remarks

■ Let $\underline{\pi}=\left(\underline{\pi}_{0}, \underline{\pi}_{1}, \ldots, \underline{\pi}_{b_{1}}\right)$ be the concatenated steady-state probability (row) vector having a total of $b_{1}+1$ levels.

- To determine $\underline{\pi}_{m}$ for $m \geq 0$, we need to solve $\underline{\tilde{0}}=\underline{\pi} Q$ where $Q$ is the ( $n_{1}+b_{1} n_{2}$ )-dimensioned infinitesimal generator of the process and $\underline{\tilde{0}}=\left(\underline{0}_{n_{1}}, \underline{0}_{n_{2}}, \ldots, \underline{0}_{n_{2}}\right)$ is an appropriately partitioned row vector (having a total of $b_{1}+1$ levels) such that $\underline{0}_{n_{i}}$ denotes a $1 \times n_{i}$ row vector of zeros.
■ $Q$ is block-structured as a level-dependent QBD process with blocks $Q_{m, j}$ containing all transitions where $X_{1}$ changes from $m$ to $j$.

|  | 0 | 1 | 2 | $b_{1}-2$ | $b_{1}-1$ | $b_{1}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | $\left(Q_{0,0}\right.$ | $Q_{0,1}$ | 0 | 0 | 0 | 0 |
| 1 | $Q_{1,0}$ | $Q_{1,1}$ | $Q_{1,2}$ | 0 | 0 | 0 |
| $Q=2$ | 0 | $Q_{2,1}$ | $Q_{2,2}$ | 0 | 0 | 0 |
| $:$ |  |  |  | : |  | : |
| $b_{1}-2$ | 0 | 0 | 0 | $Q_{b_{1}-2, b_{1}-2}$ | $Q_{b_{1}-2, b_{1}-1}$ | 0 |
| $b_{1}-1$ | 0 | 0 | 0 | $Q_{b_{1}-1, b_{1}-2}$ | $Q_{b_{1}-1, b_{1}-1}$ | $Q_{b_{1}-1, b_{1}}$ |
| $b_{1}$ | (0 | 0 | 0 | 0 | $Q_{b_{1}, b_{1}-1}$ | $Q_{b_{1}, b_{1}}$ |

## Building $Q$

## On a 2-class

Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis

■ Note that $Q_{1,2}=Q_{2,3}=\cdots=Q_{b_{1}-1, b_{1}}=\lambda_{1} I_{n_{2}}$ where $I_{n_{2}}$ denotes the $n_{2} \times n_{2}$ identity matrix.

- Define $\lambda=\lambda_{1}+\lambda_{2}$.
- Define $\underline{e}_{i, j}$ to be a row vector of length $i$, with 1 as the $j^{t h}$ entry and zeros everywhere else.
- $\otimes$ denotes the Kronecker product operator, $\delta_{i, j}$ denotes the Kronecker delta function, and the prime symbol, ${ }^{\prime}$, denotes vector transpose.
- Define $\underline{v}=\left(v_{1}, v_{2}\right), V=\operatorname{diag}(\underline{v}), V_{1}=v_{1} \underline{e}_{2,1}^{\prime} \underline{e}_{2,2}, V_{2}=v_{2} \underline{e}_{2,2}^{\prime} \underline{e}_{2,1}$.
- Some select blocks of $Q$ are as follows:


## Building $Q$

On a 2-class
Polling
Model...
Kevin
Granville \&
Steve Drekic

Introduction
and
Preliminaries
Determination
of the
Steady-state
Probabilities
Determination
of the
Waiting Time
Distribution
Numerical
Analysis
Concluding
Remarks

$$
\begin{aligned}
& \Gamma_{j}=\left[\begin{array}{cc}
(j-1) \alpha_{2} I_{k_{2} s_{2}}+U_{2} & \underline{e}_{k_{2}, k_{2}}^{\prime} \frac{e_{2,2}}{j \alpha_{2} I_{2}} \\
\mathbf{0}
\end{array}\right], \quad \underline{S}_{0,2}^{\prime}= \begin{cases}0 & \text { if } k_{i}=1, \\
{\left[\begin{array}{cc}
\underline{0}_{k_{i}-1}^{\prime} & I_{k_{i}-1} \\
0 & \underline{0}_{k_{i}-1}
\end{array}\right] \otimes \underline{s}_{0, i}^{\prime} \underline{\beta}_{i}} & \text { if } k_{i} \geq 2,\end{cases} \\
& \Delta_{j}=\left[\begin{array}{cc}
-I_{k_{2}} \otimes\left(\left(\lambda-\lambda_{2} \delta_{j, b_{2}}+(j-1) \alpha_{2}\right) I_{s_{2}}-S_{2}\right) & 0 \\
\underline{e}_{2,1}^{\prime} \underline{e}_{k_{2}, 1} \otimes v_{1} \underline{\beta}_{2} & -\left(\left(\lambda-\lambda_{2} \delta_{j, b_{2}}+j \alpha_{2}\right) I_{2}+V-v_{2}\right)
\end{array}\right],
\end{aligned}
$$

## Building $Q$

## On a 2-class Polling Model...

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries

## Determination

 of the Steady-state Probabilities
## Determination

## of the

Waiting Time Distribution

## Numerical

Analysis

## Concluding

 Remarks$$
\begin{aligned}
& 0 \begin{array}{llllll}
1 & 2 & b_{2}-1 & b_{2}
\end{array} \\
& {\left[\begin{array}{cc}
{\left[\lambda_{2} I_{k_{1} s_{1}}, \underline{0}_{\left.k_{1} s_{1} \underline{0}_{k_{2} s_{2}}^{\prime}\right]}^{0}\right.} & \mathbf{0} \\
\lambda_{2} I_{2}
\end{array}\right] \quad 0 \quad \cdots \quad 0 \quad 0} \\
& C_{i, 1} \quad \lambda_{2} I_{z} \\
& 0 \quad 0 \\
& C_{i, 2} \\
& 00 \\
& \begin{array}{l}
0 \\
0
\end{array} \\
& \left.\begin{array}{cc}
C_{i, b_{2}-1} & \lambda_{2} I_{z_{2}} \\
B_{b_{2}} & C_{i, b_{2}}
\end{array}\right)
\end{aligned}
$$

$$
B_{j}=\left[\begin{array}{cc}
j \alpha_{2} / l_{k s_{1}} & \mathbf{0} \\
\mathbf{0} & \Gamma_{j}
\end{array}\right],
$$

## Building $Q$

## On a 2-class

 Polling Model...
## Kevin

## Granville \&

 Steve DrekicIntroduction and
Preliminaries

## Determination

 of the Steady-state ProbabilitiesDetermination of the Waiting Time Distribution

## Numerical

 Analysis
## Concluding

 Remarks$$
\begin{aligned}
& \zeta_{x, i, j}=-I_{k_{x}} \otimes\left(\left(\lambda-\lambda_{1} \delta_{i, b_{1}}-\lambda_{2} \delta_{j, b_{2}}+\left(i-\delta_{x, 1}\right) \alpha_{1}+\left(j-\delta_{x, 2}\right) \alpha_{2}\right) I_{s_{x}}-S_{x}\right),
\end{aligned}
$$

## Building $Q$



## Calculating $\underline{\pi}$

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries

Determination of the Waiting Time Distribution

■ Level-dependent QBD processes are well-studied in the literature, and it is possible to develop a computational procedure for calculating the steady-state probabilities associated with our model.

- From $\underline{\tilde{0}}=\underline{\pi} Q$, the equilibrium equations in block form are obtained:

$$
\begin{align*}
& \underline{0}_{n_{1}}=\underline{\pi}_{0} Q_{0,0}+\underline{\pi}_{1} Q_{1,0}  \tag{3}\\
& \underline{0}_{n_{2}}=\underline{\pi}_{0} Q_{0,1}+\underline{\pi}_{1} Q_{1,1}+\underline{\pi}_{2} Q_{2,1},  \tag{4}\\
& \underline{0}_{n_{2}}=\lambda_{1} \underline{\pi}_{m-1}+\underline{\pi}_{m} Q_{m, m}+\underline{\pi}_{m+1} Q_{m+1, m}, m=2,3, \ldots, b_{1}-1,  \tag{5}\\
& \underline{0}_{n_{2}}=\lambda_{1} \underline{\pi}_{b_{1}-1}+\underline{\pi}_{b_{1}} Q_{b_{1}, b_{1}} . \tag{6}
\end{align*}
$$

## Calculating $\underline{\pi}$

- Solving equations (4) through (6) inductively yields

$$
\begin{equation*}
\underline{\pi}_{m}=\underline{\pi}_{0} \prod_{j=1}^{m} \mathcal{S}_{j}, m=1,2, \ldots, b_{1} \tag{7}
\end{equation*}
$$

where the set of matrices $\left\{\mathcal{S}_{j} ; j=1,2, \ldots, b_{1}\right\}$ satisfy the recursive relation

$$
\mathcal{S}_{j}=-\lambda_{1}\left(Q_{j, j}+\mathcal{S}_{j+1} Q_{j+1, j}\right)^{-1}, j=2,3, \ldots, b_{1}-1
$$

with

$$
\mathcal{S}_{b_{1}}=-\lambda_{1} Q_{b_{1}, b_{1}}^{-1} \quad \text { and } \quad \mathcal{S}_{1}=-Q_{0,1}\left(Q_{1,1}+\mathcal{S}_{2} Q_{2,1}\right)^{-1}
$$

- Defining $\mathcal{S}_{0}=Q_{0,0}+\mathcal{S}_{1} Q_{1,0}$, equation (3) becomes

$$
\begin{equation*}
\underline{\pi}_{0} \mathcal{S}_{0}=\underline{0}_{n_{1}} . \tag{8}
\end{equation*}
$$

## Calculating $\underline{\pi}$

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries

Determination of the Waiting Time Distribution

■ Since all probabilities sum to 1 , we must have that

$$
\begin{equation*}
\underline{\pi}_{0} \underline{e}_{n_{1}}^{\prime}+\underline{\pi}_{0} \mathcal{S}_{1} \underline{e}_{n_{2}}^{\prime}+\underline{\pi}_{0} \mathcal{S}_{1} \mathcal{S}_{2} \underline{e}_{n_{2}}^{\prime}+\cdots+\underline{\pi}_{0} \mathcal{S}_{1} \mathcal{S}_{2} \cdots \mathcal{S}_{b_{1}} \underline{e}_{n_{2}}^{\prime}=1 \tag{9}
\end{equation*}
$$

- Factoring out $\underline{\pi}_{0}$ and defining the column vector

$$
\underline{u}^{\prime}=\underline{e}_{n_{1}}^{\prime}+\sum_{m=1}^{b_{1}} \prod_{j=1}^{m} \mathcal{S}_{j} \underline{e}_{n_{2}}^{\prime}
$$

equations (8) and (9) give rise to the following system of linear equations which must be solved to determine $\underline{\pi}_{0}$ :

$$
\underline{\pi}_{0}\left[\begin{array}{ll}
\mathcal{S}_{0} & \underline{u}^{\prime} \tag{10}
\end{array}\right]=\left(\underline{0}_{n_{1}}, 1\right)
$$

■ In equation (10), ( $\left.\underline{0}_{n_{1}}, 1\right)$ represents the concatenated row vector of size $n_{1}+1$.

## Calculating $\underline{\pi}$

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis

- Once $\underline{\pi}_{0}$ is determined, $\underline{\pi}_{m}, m \geq 1$, is obtained via equation (7).
- Having calculated the steady-state probabilities, the blocking probabilities for each class can be defined:

$$
\begin{aligned}
& P_{b_{1}, \bullet}=\sum_{j=0}^{b_{2}} P_{b_{1}, j} \\
& P_{\bullet, b_{2}}=\sum_{m=0}^{b_{1}} P_{m, b_{2}}
\end{aligned}
$$

- These correspond to the probabilities of a class-1 or class-2 customer being turned away at entry (and subsequently lost) due to their class queue being full.
- These values are particularly useful in selecting buffer sizes $b_{1}$ and $b_{2}$ so as to ensure negligible blocking probabilities are obtained for both queues.


## Nominal Waiting Time

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution Remarks

■ For $i=1,2$, let $W_{i}$ represent the duration of time from the (successful) arrival of an arbitrary class- $i$ customer to the system until the server is reached, referred to as the nominal class- $i$ waiting time.

- Without loss of generality, we focus our analysis only on $W_{1}$ as the characteristics of the two queues are essentially indifferent.
- Define the modified steady-state probabilities

$$
\phi_{0,0, l, 0}=\frac{\pi_{0,0, l, 0}}{1-P_{b_{1}, \bullet}} \text { and } \phi_{m, n, l, y}=\frac{\pi_{m, n, l, y}}{1-P_{b_{1}, \bullet}}
$$

where $m=1,2, \ldots, b_{1}-1, n=1,2, \ldots, b_{2}$, and the components $I$ and $y$ are as defined in equations (1) and (2), respectively.

## Nominal Waiting Time

- Several row vectors are required in the subsequent analysis such as:

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis
Concluding Remarks

$$
\begin{gathered}
\underline{\phi}_{0, n}=\frac{\underline{\underline{I}}_{0, n}}{1-P_{b_{1}, \bullet}}, 1 \leq n \leq b_{2}, \\
\underline{\phi}_{m, 0}=\frac{\underline{\pi}_{m, 0}}{1-P_{b_{1}, \bullet}}, 1 \leq m \leq b_{1}-1, \\
\underline{\phi}_{m, n}=\frac{\underline{\pi}_{m, n}}{1-P_{b_{1}, \bullet}}, 1 \leq m \leq b_{1}-1,1 \leq n \leq b_{2}
\end{gathered}
$$

- Furthermore, let

$$
\underline{\phi}_{0}=\left(\phi_{0,0, k_{1}+k_{2}+1,0}, \phi_{0,0, k_{1}+k_{2}+2,0}, \underline{\phi}_{0,1}, \underline{\phi}_{0,2}, \ldots, \underline{\phi}_{0, b_{2}}\right)
$$

and

$$
\underline{\phi}_{m}=\left(\underline{\phi}_{m, 0}, \underline{\phi}_{m, 1}, \ldots, \underline{\phi}_{m, b_{2}}\right), m=1,2, \ldots, b_{1}-1 .
$$

## Nominal Waiting Time

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis

Concluding Remarks

- By constructing

$$
\begin{equation*}
\underline{\Phi}=\left(\underline{\phi}_{b_{1}-1}, \underline{\phi}_{b_{1}-2}, \ldots, \underline{\phi}_{1}, \underline{\phi}_{0}\right) \tag{11}
\end{equation*}
$$

to be the concatenated row vector of dimension

$$
\begin{equation*}
\ell=\left(b_{1}-1\right) n_{2}+n_{1} \tag{12}
\end{equation*}
$$

we note that $\underline{\Phi} \underline{e}_{\ell}^{\prime}=1$.

- Upon successful entry into one of the $\ell$ possible busy states, the PASTA property ensures that our Poisson-arriving class-1 customer finds the system in state ( $m, n, l, y$ ) with probability $\phi_{m, n, l, y}$.
- For now, we assume that the target class-1 customer is not subject to reneging.
- While waiting in the class-1 queue, the number of customers in the class-2 queue potentially changes, as well as the indicator on the server which identifies how many customers have completed service in the active serving queue.


## Nominal Waiting Time

On a 2-class
Polling Model...

Kevin Granville \& Steve Drekic

Introduction and Preliminaries

Determination of the Steady-state Probabilities
Determination of the Waiting Time Distribution

Numerical Analysis

Concluding Remarks

- As the number of customers in the class-1 queue changes, the ones arriving later have no impact on the waiting time of the target class-1 customer.
- If we effectively think of the arrival rate for the class- 1 queue to be equal to 0 , the distribution of $W_{1}$ can be modelled as the distribution of the time to absorption in a Markov chain with infinitesimal generator of the form

$$
\left[\begin{array}{cc}
\mathcal{R} & -\mathcal{R} \underline{e}_{\ell}^{\prime} \\
\underline{0}_{\ell} & 0
\end{array}\right],
$$

where

|  | $b_{1}-1$ | $b_{1}-2$ | $b_{1}-3$ | 2 | 1 | 0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $b_{1}-1$ | $\widetilde{Q}^{\widetilde{Q}_{b_{1}-1, b_{1}-1}}$ | $Q_{b_{1}-1, b_{1}-2}$ | 0 | 0 | 0 | 0 |
| $b_{1}-2$ | 0 | $\widetilde{Q}_{b_{1}-2, b_{1}-2}$ | $Q_{b_{1}-2, b_{1}-3}$ | 0 | 0 | 0 |
| $\mathcal{R}={ }^{\text {b }}$ ( -3 | 0 | 0 | $\widetilde{Q}_{b_{1}-3, b_{1}-3}$ | 0 | 0 | 0 |
| : | : |  |  | . | . | $\vdots$ |
| 2 | 0 | 0 | 0 | $\widetilde{Q}_{2,2}$ | $Q_{2,1}$ | 0 |
| 1 | 0 | 0 | 0 | 0 | $\widetilde{Q}_{1,1}$ | $\widetilde{Q}_{1,0}$ |
| 0 | ( 0 | 0 | 0 | 0 | 0 | $\widetilde{Q}_{0,0}$ |

## Nominal Waiting Time

## On a 2-class Polling Model...

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination
of the
Steady-state Probabilities

## Determination

 of the Waiting Time Distribution■ In equation (13), $Q_{2,1}, Q_{3,2}, \ldots, Q_{b_{1}-1, b_{1}-2}$ are the same matrices defined earlier and $\widetilde{Q}_{m, m}=Q_{m, m}+\lambda_{1} I_{n_{2}}, m=1,2, \ldots, b_{1}-1$.

- In addition,

|  | 0 | 1 | 2 | . | $b_{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | $\left(\left[\begin{array}{c}\underline{e}_{k_{1}, k_{1}}^{\prime} \underline{e}_{2,1} \otimes \underline{S}_{0,1}^{\prime} \\ \alpha_{1} I_{2}\end{array}\right]\right.$ | 0 | 0 | . | 0 |
| 1 | 0 |  | 0 | . | 0 |
| $\tilde{Q}_{1,0}=2$ | 0 | 0 | $\left[\begin{array}{l}\underline{e}_{k_{1}, k_{1}}^{\prime} e_{\underline{z}_{1}, z_{1}-1} \otimes \underline{S}_{0,1}^{\prime} \\ \alpha_{1} I_{z_{1}}\end{array}\right]$ | $\ldots$ | 0 |
| : |  | : | - | $\because$ |  |
| $b_{2}$ | 0 | 0 | 0 | $\ldots$ |  |

## Nominal Waiting Time

## On a 2-class Polling <br> Model...

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination
of the
Steady-state Probabilities
Determination of the Waiting Time Distribution

Numerical Analysis

Concluding Remarks

- Also,

$$
\text { where } \tilde{\Delta}_{i}=\Delta_{i}+\operatorname{diag}\left(\lambda_{1} I_{k_{2} s_{2}}, \lambda_{1} I_{2}-V_{2}\right)
$$

$$
\begin{aligned}
& 0 \begin{array}{ccccc}
1 & 2 & \ldots & b_{2}-1 & b_{2}
\end{array}
\end{aligned}
$$

## Actual Delay Distribution

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis

Concluding Remarks

- The time to absorption distribution of such a Markov chain has received extensive attention in the literature, and it is well-known that the cumulative distribution function of $W_{1}$, denoted by $F_{1}(\omega)$, is given by

$$
F_{1}(\omega)=1-\Phi \exp \{\mathcal{R} \omega\} \underline{e}_{\ell}^{\prime}, \omega \geq 0
$$

which is of phase-type form.

- To incorporate the reneging behaviour of our target class-1 customer, define $W_{1}^{*}$ to be the actual class-1 delay (i.e., the arriving class- 1 customer's total time spent in the system prior to successfully entering service).
■ Clearly, $G_{1}(\omega)=\operatorname{Pr}\left(W_{1}^{*} \leq \omega\right)=\operatorname{Pr}\left(W_{1} \leq \omega \mid W_{1} \leq R_{1}\right)$, where $R_{1}$ denotes an exponentially distributed random variable, independent of $W_{1}$, with rate $\alpha_{1}$.
- Making use of fundamental matrix algebraic techniques, the following expressions for $G_{1}(\omega)$ and the moments of $W_{1}^{*}$ are obtained:


## Actual Delay Distribution

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

## Determination

 of the Waiting Time Distribution$$
\begin{aligned}
G_{1}(\omega) & =1-\operatorname{Pr}\left(W_{1}>\omega \mid W_{1} \leq R_{1}\right) \\
& =1-\frac{\operatorname{Pr}\left(\omega<W_{1} \leq R_{1}\right)}{\operatorname{Pr}\left(W_{1} \leq R_{1}\right)} \\
& =1-\frac{\int_{\omega}^{\infty} \operatorname{Pr}\left(W_{1}>\omega\right) \alpha_{1} e^{-\alpha_{1} x} d x-\int_{\omega}^{\infty} \operatorname{Pr}\left(W_{1}>x\right) \alpha_{1} e^{-\alpha_{1} x} d x}{1-\int_{0}^{\infty} \operatorname{Pr}\left(W_{1}>x\right) \alpha_{1} e^{-\alpha_{1} x} d x} \\
& =1-\frac{\Phi\left[I_{\ell}-\alpha_{1}\left(\alpha_{1} I_{\ell}-\mathcal{R}\right)^{-1}\right] \exp \{\mathcal{R} \omega\} \underline{e}_{\ell}^{\prime} e^{-\alpha_{1} \omega}}{1-\alpha_{1} \Phi\left(\alpha_{1} I_{\ell}-\mathcal{R}\right)^{-1} \underline{e}_{\ell}^{\prime}}, \omega \geq 0,
\end{aligned}
$$

and

$$
\mathrm{E}\left[W_{1}^{* r}\right]=\frac{r!\Phi\left[I_{\ell}-\alpha_{1}\left(\alpha_{1} I_{\ell}-\mathcal{R}\right)^{-1}\right]\left(\alpha_{1} I_{\ell}-\mathcal{R}\right)^{-r} \underline{e}_{\ell}^{\prime}}{1-\alpha_{1} \Phi\left(\alpha_{1} I_{\ell}-\mathcal{R}\right)^{-1} \underline{e}_{\ell}^{\prime}}, r=1,2, \ldots
$$

## Total Time Spent Waiting In The System

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination

- We investigate the selection of parameters $k_{1}$ and $k_{2}$ in order to optimize the overall system, by way of minimizing a specific cost function.
- The total time a class- 1 customer actually spends waiting in the system is

$$
W_{1}^{\#}=\min \left\{W_{1}, R_{1}\right\} \sim P H\left(\Phi_{1}, \mathcal{R}_{1}-\alpha_{1} I_{\ell_{1}}\right),
$$

where $\Phi_{1}, \ell_{1}$, and $\mathcal{R}_{1}$ are given by equations (11), (12), and (13), respectively.

- Parameters $\Phi_{2}, \ell_{2}$, and $\mathcal{R}_{2}$ can be obtained in an analogous fashion to be used in the characterization of the distribution of $W_{2}^{\#}=\min \left\{W_{2}, R_{2}\right\}$.


## The Cost Function

■ We generalize the cost function of Borst, Boxma \& Levy (1995) to get

$$
\text { Cost }=\text { Cost }_{1}+\text { Cost }_{2},
$$

where

$$
\operatorname{Cost}_{i}=c_{i} \lambda_{i} \mathrm{E}\left[W_{i}^{\#}\right]+r_{i} \lambda_{i} \operatorname{Pr}\left(R_{i}<W_{i}\right)
$$

and cost parameters $c_{i}$ and $r_{i}$ are non-negative constants.

- It is straightforward to obtain:

$$
\begin{aligned}
\mathrm{E}\left[W_{1}^{\#}\right] & =\underline{\Phi}\left(\alpha_{1} I_{\ell_{1}}-\mathcal{R}_{1}\right)^{-1} \underline{e}_{\ell_{1}}^{\prime} \\
\operatorname{Pr}\left(R_{1}<W_{1}\right) & =\alpha_{1} \Phi\left(\alpha_{1} I_{\ell_{1}}-\mathcal{R}_{1}\right)^{-1} \underline{e}_{\ell_{1}}^{\prime}=\alpha_{1} \mathrm{E}\left[W_{1}^{\#}\right] .
\end{aligned}
$$

## Optimization Problem

On a 2-class
Polling
Model...
Kevin
Granville \& Steve Drekic

Introduction and
Preliminaries
Determination of the Steady-state Probabilities

Determination of the Waiting Time Distribution

Numerical Analysis
Concluding Remarks

- We consider a constraint on the total number of services in a cycle - namely, $k_{1}+k_{2} \leq 12$.
- Three possible reneging rates were used for both classes:

$$
\alpha_{i} \in\{0.025,0.05,0.25\} .
$$

- Three service time distributions were considered: Exponential [Exp], Hyperexponential $\left[\mathrm{H}_{2}\right]$, and Erlang [ $\mathrm{E}_{3}$ ].
- Each distribution has the same mean, but the $\mathrm{H}_{2}$ distribution has 1000 times the variance of the Exp distribution, which has 3 times the variance of the $E_{3}$ distribution.
■ Case 1: arrival rates $\lambda_{1}=\lambda_{2}=0.75$, switchover rates $v_{1}=v_{2}=1 / 0.1$, and mean service times of 0.9 for class 1 and 0.1 for class 2 .
■ Case 2: arrival rates $\lambda_{1}=0.5, \lambda_{2}=0.25$, switchover rates $v_{1}=1 / 0.1$, $v_{2}=1 / 0.2$, and mean service times of 1 for both classes.
■ In both cases, we set buffer sizes of $b_{1}=b_{2}=20$.


## Numerical Analysis



| Reneging Rates |  | (Exp, Exp) |  | Service Time Distributions |  |  |  | (Exp, Exp) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\alpha_{1}$ | $\alpha_{2}$ | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost |
| 0.025 | 0.025 | $(3,9)$ | 4.3398 | $(3,9)$ | 6.2866 | $(3,9)$ | 4.3281 | $(3,9)$ | 6.9361 |
|  | 0.05 | $(4,8)$ | 4.2581 | $(3,9)$ | 5.9977 | $(4,8)$ | 4.2468 | $(2,10)$ | 7.7429 |
|  | 0.25 | $(7,5)$ | 3.7352 | $(9,3)$ | 4.8325 | $(7,5)$ | 3.7269 | $(2,10)$ | 11.9875 |
| 0.05 | 0.025 | $(3,9)$ | 3.6482 | $(3,9)$ | 5.2460 | $(3,9)$ | 3.6386 | $(3,9)$ | 7.1422 |
|  | 0.05 | $(3,9)$ | 3.5947 | $(3,9)$ | 4.9824 | $(3,9)$ | 3.5855 | $(2,10)$ | 7.8847 |
|  | 0.25 | $(6,6)$ | 3.2543 | $(6,6)$ | 4.0882 | $(6,6)$ | 3.2470 | $(1,11)$ | 11.9519 |
| 0.25 | 0.025 | $(2,10)$ | 2.1520 | $(2,10)$ | 3.0167 | $(2,10)$ | 2.1464 | $(3,9)$ | 9.0264 |
|  | 0.05 | $(2,10)$ | 2.1334 | $(2,10)$ | 2.8169 | $(2,10)$ | 2.1279 | $(3,9)$ | 9.7272 |
|  | 0.25 | $(2,10)$ | 2.0230 | $(2,10)$ | 2.2667 | $(2,10)$ | 2.0183 | $(1,11)$ | 13.3357 |
|  |  | $\left(\mathrm{H}_{2}, \mathrm{Exp}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{H}_{2}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{E}_{3}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{H}_{2}\right)$ |  |
| $\alpha_{1}$ | $\alpha_{2}$ | $\left(k_{1}, k_{2}\right)$ | Cost | ( $k_{1}, k_{2}$ ) | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost |
| 0.025 | 0.025 | $(4,8)$ | 20.3486 | $(3,9)$ | 21.7547 | $(4,8)$ | 20.3441 | $(3,9)$ | 35.8657 |
|  | 0.05 | $(5,7)$ | 18.2711 | $(4,8)$ | 19.5065 | $(5,7)$ | 18.2666 | $(3,9)$ | 36.3322 |
|  | 0.25 | $(8,4)$ | 14.6158 | $(9,3)$ | 15.5758 | $(8,4)$ | 14.6114 | $(2,10)$ | 33.8738 |
| 0.05 | 0.025 | $(3,9)$ | 16.4205 | $(2,10)$ | 17.4343 | $(3,9)$ | 16.4171 | $(2,10)$ | 34.1935 |
|  | 0.05 | $(4,8)$ | 14.3710 | $(3,9)$ | 15.2235 | $(4,8)$ | 14.3676 | $(3,9)$ | 34.6786 |
|  | 0.25 | $(7,5)$ | 10.7526 | $(8,4)$ | 11.3615 | $(7,5)$ | 10.7490 | $(2,10)$ | 32.2619 |
| 0.25 | 0.025 | $(1,11)$ | 8.8681 | $(1,11)$ | 9.4376 | $(1,11)$ | 8.8681 | $(3,9)$ | 27.1216 |
|  | 0.05 | $(1,11)$ | 6.9769 | $(1,11)$ | 7.3890 | $(1,11)$ | 6.9756 | $(3,9)$ | 27.6632 |
|  | 0.25 | $(6,6)$ | 3.5293 | $(5,7)$ | 3.7245 | $(6,6)$ | 3.5263 | $(2,10)$ | 25.4136 |
| $r_{i}$ |  | $r_{1}=1, r_{2}=0.5$ |  |  |  |  |  | $r_{1}=r_{2}=40$ |  |

Table 1: Optimal $\left(k_{1}, k_{2}\right)$ and minimum cost under Case 1 with $c_{1}=2, c_{2}=1$, and $r_{1}=1$, $r_{2}=0.5$ or $r_{1}=r_{2}=40 .\left[\lambda_{1}=\lambda_{2}=0.75, v_{1}=v_{2}=1 / 0.1\right.$, service times of $\left.0.9 \& 0.1\right]$

## Numerical Analysis



| Reneging Rates |  | (Exp, Exp) |  | Service Time Distributions$\left(\operatorname{Exp}, \mathrm{H}_{2}\right)$$\left(\operatorname{Exp}, \mathrm{E}_{3}\right)$ |  |  |  | (Exp, Exp) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\alpha_{1}$ | $\alpha_{2}$ | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost |
| 0.025 | 0.025 | $(10,2)$ | 2.6649 | $(10,2)$ | 7.8090 | $(10,2)$ | 2.4761 | $(10,2)$ | 4.3972 |
|  | 0.05 | $(11,1)$ | 2.4083 | $(11,1)$ | 6.8184 | $(11,1)$ | 2.2577 | $(9,3)$ | 4.6890 |
|  | 0.25 | $(11,1)$ | 1.8357 | $(11,1)$ | 5.1854 | $(11,1)$ | 1.7432 | $(8,4)$ | 5.9054 |
| 0.05 | 0.025 | $(10,2)$ | 2.3812 | $(9,3)$ | 5.6575 | $(10,2)$ | 2.2247 | $(10,2)$ | 4.6660 |
|  | 0.05 | $(10,2)$ | 2.2030 | $(10,2)$ | 4.8100 | $(10,2)$ | 2.0669 | $(10,2)$ | 4.9625 |
|  | 0.25 | $(11,1)$ | 1.6934 | $(11,1)$ | 3.5890 | $(11,1)$ | 1.6127 | $(8,4)$ | 6.1953 |
| 0.25 | 0.025 | $(6,6)$ | 1.5047 | $(5,7)$ | 2.9877 | $(6,6)$ | 1.4353 | $(11,1)$ | 6.4562 |
|  | 0.05 | $(7,5)$ | 1.4550 | $(6,6)$ | 2.2615 | $(7,5)$ | 1.3907 | $(10,2)$ | 6.7245 |
|  | 0.25 | $(11,1)$ | 1.2323 | $(10,2)$ | 1.5454 | $(11,1)$ | 1.1879 | $(8,4)$ | 7.8861 |
|  |  | $\left(\mathrm{H}_{2}, \mathrm{Exp}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{H}_{2}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{E}_{3}\right)$ |  | $\left(\mathrm{H}_{2}, \mathrm{H}_{2}\right)$ |  |
| $\alpha_{1}$ | $\alpha_{2}$ | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost | $\left(k_{1}, k_{2}\right)$ | Cost |
| 0.025 | 0.025 | $(11,1)$ | 11.9085 | $(10,2)$ | 15.6149 | $(11,1)$ | 11.8306 | $(10,2)$ | 24.8349 |
|  | 0.05 | $(11,1)$ | 10.6576 | $(11,1)$ | 13.9119 | $(11,1)$ | 10.5843 | $(10,2)$ | 23.2350 |
|  | 0.25 | $(11,1)$ | 9.5631 | $(11,1)$ | 12.3096 | $(11,1)$ | 9.5085 | $(9,3)$ | 21.9772 |
| 0.05 | 0.025 | $(10,2)$ | 8.2620 | $(9,3)$ | 10.5575 | $(10,2)$ | 8.1927 | $(9,3)$ | 20.6821 |
|  | 0.05 | $(11,1)$ | 7.0367 | $(10,2)$ | 8.9056 | $(11,1)$ | 6.9730 | $(9,3)$ | 19.1262 |
|  | 0.25 | $(11,1)$ | 5.9644 | $(11,1)$ | 7.4260 | $(11,1)$ | 5.9166 | $(8,4)$ | 17.9732 |
| 0.25 | 0.025 | $(4,8)$ | 3.9129 | $(4,8)$ | 5.0239 | $(4,8)$ | 3.8846 | $(8,4)$ | 15.7298 |
|  | 0.05 | $(9,3)$ | 2.8207 | $(6,6)$ | 3.4547 | $(9,3)$ | 2.7818 | $(8,4)$ | 14.2528 |
|  | 0.25 | $(11,1)$ | 1.8533 | $(9,3)$ | 2.1237 | $(11,1)$ | 1.8198 | $(7,5)$ | 13.2301 |
| $r_{i}$ |  | $r_{1}=1, r_{2}=0.5$ |  |  |  |  |  | $r_{1}=r_{2}=40$ |  |

Table 2: Optimal $\left(k_{1}, k_{2}\right)$ and minimum cost under Case 2 with $c_{1}=2, c_{2}=1$, and $r_{1}=1$, $r_{2}=0.5$ or $r_{1}=r_{2}=40 .\left[\lambda_{1}=0.5, \lambda_{2}=0.25, v_{1}=1 / 0.1, v_{2}=1 / 0.2\right.$, service times of 1$]$

## Numerical Analysis

Determination
of the
Steady-state
Probabilities

On a 2-class Polling Model...

Kevin Granville \& Steve Drekic

Introduction and

```
Preliminaries
```

```
Preliminaries
```


## Determination

## of the

Waiting Time Distribution

## Numerical

 Analysis
## Concluding

 RemarksCase 1


Case 2


Figure 1: Plots of $k_{1}$ vs. $c_{1}$ under both cases with Exp service times, $c_{2}=2, r_{1}=r_{2}=1$, and four combinations of reneging rates.

## Numerical Analysis

On a 2-class Polling Model...

Kevin Granville \& Steve Drekic

Introduction and
Preliminaries
Determination
of the
Steady-state Probabilities

Determination

## of the

Waiting Time Distribution

## Numerical

 Analysis
## Concluding

 RemarksClass 1


Class 2


Figure 2: Plots of $G_{i}(\omega)$ vs. $\omega$ for both classes under Case 1 with $\alpha_{1}=0.025, \alpha_{2}=0.25$, and either Exp or $\mathrm{H}_{2}$ service time distributions, at optimal $k_{i}$ 's from Table 1.

## Concluding Remarks

On a 2-class
Polling
Model...
Kevin Granville \& Steve Drekic

Introduction and Preliminaries

Determination of the Steady-state Probabilities
Determination of the Waiting Time Distribution
Numerical Analysis
Concluding Remarks

■ We modelled an $M / P H / 1$-type polling model with 2 classes and exponential switchover/reneging times operating under a $k_{-}$-limited service discipline.

- We were able to obtain the steady-state probabilities as well as the nominal waiting time and actual delay distributions for each class.
■ Under a variety of scenarios, we found optimal $\left(k_{1}, k_{2}\right)$ values that minimize (subject to a particular constraint) a defined cost function depending on the expected time waiting in the system and the probability of reneging.


## Future Extensions

On a 2-class Polling Model...

## Kevin

 Granville \& Steve Drekic- Queue length dependent reneging rates.
- Phase-type renewal process for arrivals.
- (2-dimensional) phase-type reneging time distributions.
- Different service disciplines.
- Multiple servers.
- A third class of customers.
On a 2-class


## Kevin

Granville \&
Steve Drekic

anc

Preliminaries

Determination

of the
Steady-state
Probabilities
Determination
of the
Waiting Time
Distribution
Numerical
Analysis

## Building $Q$

On a 2-class
Polling
Model...
Kevin
Granville \&
Steve Drekic

Introduction
and
Preliminaries
Determination
of the
Steady-state
Probabilities
Determination
of the
Waiting Time
Distribution
Numerical
Analysis
Concluding
Remarks

$$
\begin{aligned}
& 1 \quad 2 \\
& 2 \quad \cdots \quad b_{2}
\end{aligned}
$$

